

A Bayesian Approach for the Rainfall-Runoff Problem in Multiple Basins

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Introduction: Rainfall-Runoff relationship

Proposed approach: Rainfall-Runoff in multiple basins

Rainfall-Runoff model

Let Y_t be the basin **runoff** and X_t the basin **rainfall** at time t:

 $Y_t \sim p(Y_t | \mu_t, \phi_t), \quad t = 1, 2, ...$ $g(\mu_t) = f_1(\alpha_t, E_t)$ $E_t = f_2(E_{t-1}, \ldots, E_0, X_t)$

- $-p(Y_t|\mu_t, \phi_t)$ defined in \mathbb{R}^+ : Gamma, log-normal, etc;
- $-\alpha_t$ denotes a trend and E_t denotes the total effect of rainfall at time t;
- $-g(\cdot), f_1(\cdot)$ and $f_2(\cdot)$ are known non-linear functions describing the dynamics of the hydrological process.

Rainfall Effect on Runoff: A transfer function

Following Migon and Monteiro (1997):

 $E_t = \rho_t E_{t-1} + \gamma_t X_t$

Multivariate Rainfall-Runoff model

Let Y_t^m be the runoff at time t from basin m, t = 1, ..., T; m = 1, ..., M. The joint distribution of $Y_t = (Y_t^1, \dots, Y_t^M)'$ can be expressed, for example, as:

 $p(\boldsymbol{Y}_t|\boldsymbol{\theta}) = p(\boldsymbol{Y}_t^M|\boldsymbol{Y}_t^{M-1},\boldsymbol{\theta})p(\boldsymbol{Y}_t^{M-1}|\boldsymbol{Y}_t^{M-2},\boldsymbol{\theta})\cdots p(\boldsymbol{Y}_t^2|\boldsymbol{Y}_t^1,\boldsymbol{\theta})p(\boldsymbol{Y}_t^1,\boldsymbol{\theta}),$

where $p(Y_t^m)$ is the (conditional) distribution of Y_t^m .

Rainfall Effect

Let X_t^m be the rainfall at time t from basin m. Let $m = A, B, B \subset A$, then $Y_t = (Y_t^A, Y_t^B)'$ and $X_t = (X_t^A, X_t^B)'$ are the time series of both basins. Let $X_t^{A|B} = X_t^A - X_t^B$ be the rainfall at an area that belongs to basin A but not to basin B, then

$E_{t} = \rho_{t} E_{t-1} + [1 - \exp(-\kappa_{t} X_{t})] [\vartheta_{t} - (\alpha_{t} + \rho_{t} E_{t-1})].$

– Parameters interpretation:

- Some particular cases:

 α is the basic level or **stream flow**, ρ is the recharge factor or permanence rate of the rainfall effect and γ is the velocity of response to precipitation, related to soil saturation.

 $\gamma_t = \gamma$ $\gamma_t = \gamma_{t-1} + \delta_t$ $\gamma_t = \gamma + \delta_t; \quad \gamma \sim N(a, b)$

Basin Rainfall

Let $\{X_t(s), s \in B \subset \mathbb{R}^2, t = 1, 2, \dots\}$ be a stochastic process at discrete time t and over spatial domain B. In particular, $X_t(s)$ here represents the level of rainfall at time t and location s. The basin rainfall at time t is

 $X_t = |B|^{-1} \int_{B} X_t(s) ds,$

where |B| is the basin area. Following Sansó and Guenni (2000), $X_t(s)$ can be:

$$X_{t}(s_{i}) = \begin{cases} w_{t}(s_{i})^{\beta} & \text{if } w_{t}(s_{i}) > 0, \quad s_{i} \in B, \quad i = 1, \dots, S, \\ 0 & \text{if } w_{t}(s_{i}) \leq 0, \end{cases}$$
$$w_{t}(s) = \theta_{t}f(s) + Z_{t}(s) + \epsilon_{t}(s), \quad s = (s_{1}, \dots, s_{S})'$$
$$Z_{t}(s) \sim GP(\mathbf{0}, \sigma^{2}\varrho(||s_{1}, s_{2}||, \lambda)), \end{cases}$$

where $w_t(s_i)$ is a latent variable; $\theta_t f(s)$ is a polynomial trend; $Z_t(s)$ is a process with variance σ^2 and spatial correlation function $\rho(||s - s'||, \lambda)$ (depending on λ); and $\epsilon_t(s)$ is a random error.

 $p(\boldsymbol{Y}_t|\boldsymbol{X}_t,\boldsymbol{\theta}) = p(\boldsymbol{Y}_t^A|\boldsymbol{Y}_t^B,\boldsymbol{X}_t^{A|B},\boldsymbol{\Theta})p(\boldsymbol{Y}_t^B|\boldsymbol{X}_t^B,\boldsymbol{\Theta})$

Basins Rainfall

Following Gelfand et al. (2001) one can use Monte Carlo integration, such that:

$$X_t^m = \frac{1}{|m|} \int_m X_t(s) ds \approx \frac{1}{N_m} \sum_{i=1}^{N_m} \hat{X}_t(s_i) \quad i = 1, \dots, N_m,$$

where N_m is the number of points of a grid constructed **inside the limits** of the basin *m* and $\dot{X}_t(i)$ is the interpolated value for the s_i location of that grid.

Inference Procedure

The joint distribution of $Y = (Y_1, ..., Y_T)', X = (X_1, ..., X_T)'$ and $X_t(s) = (X_t(s_1), ..., X_t(s_S))'$ is $p(\boldsymbol{Y}, \boldsymbol{X}, \boldsymbol{X}(\boldsymbol{s}) | \boldsymbol{\Theta}) = \prod_{t=1}^{I} p(\boldsymbol{Y}_t | \boldsymbol{X}_t, \boldsymbol{X}_t(\boldsymbol{s}), \boldsymbol{\Theta}_{\boldsymbol{Y}}) p(\boldsymbol{X}_t | \boldsymbol{X}_t(\boldsymbol{s}), \boldsymbol{\Theta}_{\boldsymbol{X}}) \prod_{i=1}^{S} p(\boldsymbol{X}_t(\boldsymbol{s}_i) | \boldsymbol{\Theta}_{\boldsymbol{X}}),$

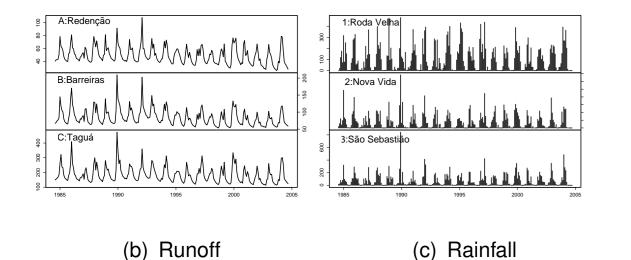
where $\Theta = (\Theta_Y, \Theta_X)$, Θ_Y are parameters in the runoff model and Θ_X are parameters in the rainfall model.

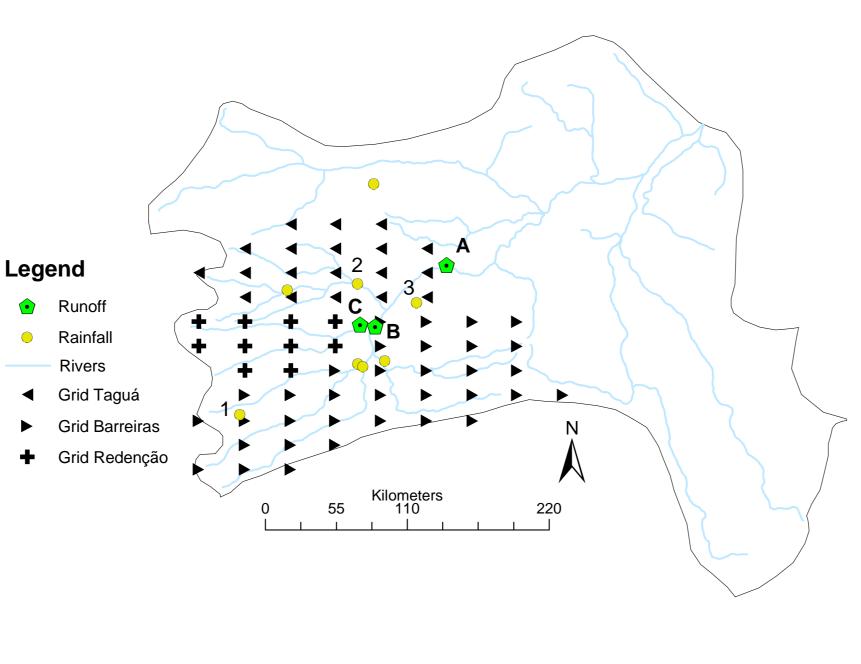
MCMC techniques: Conjugate updating backward sampling, CUBS (Ravines, Migon, & Schmidt, 2007), for Gamma transfer functions.

Application: Rio Grande Basin



(a) Location





(d) Sub-basins and Stations

Results: Multivariate transfer functions

| The model: | Ta |
|---|----|
| $Y_t^A Y_t^B \sim \text{Lognormal}(\mu_t^{A B}, \sigma_{A B}^2) t = 1, \dots, T$ | |
| $Y_t^B Y_t^C \sim \text{Lognormal}(\mu_t^{B C}, \sigma_{B C}^2)$ | |
| $Y_t^C \sim \text{Lognormal}(\mu_t^C, \sigma_C^2)$ | |
| $\mu_t^{A B} = \alpha^{A B} + \eta^{A B} Y_t^B + E_t^{A B}$ | |
| $\mu_t^{B C} = \alpha^{B C} + \eta^{B C} \Upsilon_t^C + E_t^{B C}$ | |
| $\mu_t^C = \alpha^C + E_t^C$ | |
| $E_t^m = \rho^m E_{t-1}^m + \gamma^m X_t^m m + w_t^m$ | |
| $w_t^m \sim \text{Normal}(0, W_m)$ $m = A B, B C, C;$ | |

 Table 1: Comparison with univariate models
 MSE MAE Multivariate model A|B7.025 120.684 B|C1.820 9.060 2.568 1.009 Univariate independent models 7.756 162.991 AВ 8.047 1.652 2.568 1.009

Table 2: Some posterior summaries

Figure 1: Rio Grande Basin: Location, sub-basins, monitoring sites, interpolation grids and some time series.

Data: monthly recorded series from January 1984 to September 2004, at three runoff stations and nine rainfall stations irregularly located in an area of drainage of $37522.48 \, km^2$.

Results: Spatial Interpolation

| 4 | | | |
|---|--|--|--|

| - | A B = Taguá Barreiras | | | | | | | B C = Barreiras Redenção | | | | | | C = Redenção | | | | | |
|---|-------------------------|-------|-------|-------|-------|------|----------------|----------------------------|-------|-------|-------|------|--------------|--------------|-------|-------|-------|------|--|
| _ | | mean | 25% | 50% | 75% | Ŕ | | mean | 25% | 50% | 75% | Ŕ | | mean | 25% | 50% | 75% | Ŕ | |
| | $\alpha^{A B}$ | 1.151 | 1.022 | 1.142 | 1.276 | 1.00 | $\alpha^{B C}$ | 0.587 | 0.507 | 0.581 | 0.654 | 1.01 | α^{C} | 3.543 | 3.525 | 3.544 | 3.561 | 1.01 | |
| | $\eta^{A B}$ | 0.915 | 0.887 | 0.916 | 0.945 | 1.01 | $\eta^{B C}$ | 0.989 | 0.971 | 0.990 | 1.009 | 1.01 | | | | | | | |
| | $\rho^{A B}$ | 0.510 | 0.387 | 0.533 | 0.658 | 1.00 | $\rho^{B C}$ | 0.743 | 0.692 | 0.752 | 0.806 | 1.00 | $ ho^C$ | 0.594 | 0.575 | 0.594 | 0.613 | 1.03 | |
| | $\gamma^{A B}$ | 0.025 | 0.014 | 0.024 | 0.035 | 1.00 | $\gamma^{B C}$ | 0.005 | 0.003 | 0.005 | 0.008 | 1.00 | γ^{C} | 1.645 | 1.597 | 1.645 | 1.696 | 1.00 | |
| | $W_{A B}$ | 0.003 | 0.002 | 0.003 | 0.004 | 1.01 | $W_{B C}$ | 0.003 | 0.002 | 0.002 | 0.003 | 1.00 | W_C | 0.005 | 0.005 | 0.005 | 0.006 | 1.01 | |
| | σ^2 | 0 007 | 0 006 | 0 007 | 0 009 | 1 04 | σ^2 | 0 004 | 0.003 | 0 004 | 0.005 | 1 07 | σ^2 | 0 002 | 0.002 | 0 002 | 0 003 | 1 10 | |

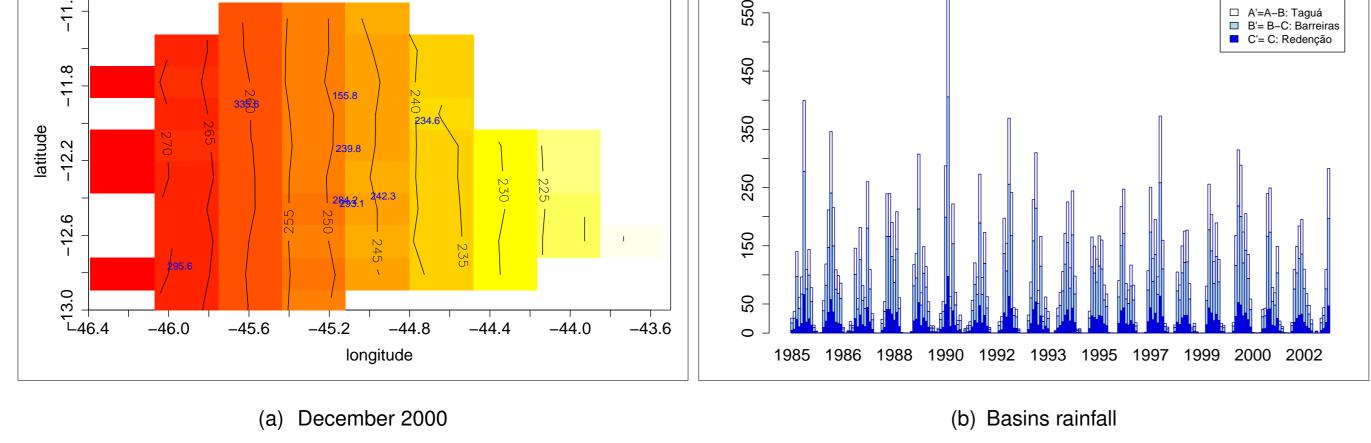


Figure 2: Posterior mean of rainfall in a specific month and the aereal rainfall for all the sub-basins

0.007 0.006 0.007 0.009 1.04 0⁻_{BlC} 0.004 0.003 0.004 0.005 1.07 0⁻_C 0.002 0.002 0.002 0.003 1.10

Final Remarks

• Main contribution: proposed approach. Modeling simultaneously rainfall and runoff taking into account the different spatial units in which they are measured.

• Features of the proposed models: its parameters have **physical interpretations** and assumptions of normality or stationarity of the time series are not needed.

• Results show that our approach is a **promising tool** for the runoff-rainfall analysis.

• An extension: hierarchical models to handle data from several basins simultaneously.

References

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